

Multiplicative Tariffs and Marginal–Sum Equations

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Multiplicative Tariff (1)

Motor car portfolio of an insurance company (exposures):

Region	Mileage in km/year		
	0-20.000	20.000-40.000	40.000- ∞
⋮ DD	600	300	100
⋮ SB	180	75	20
⋮			

Rating factors:

- ▶ Region
- ▶ Mileage

The exposure in a single cell may be too small for statistics.

Multiplicative Tariff (2)

Net premiums in percent of the overall premium level:

Region	Mileage in km/year			Factor
	0-20.000	20.000-40.000	40.000-∞	
⋮ DD	0,54	0,72	0,90	0,90
⋮ SB	0,42	0,56	0,70	0,70
⋮				
Factor	0,60	0,80	1,00	

Advantages of a multiplicative tariff:

- ▶ The number of **tariff levels** is smaller than the number of **tariff cells**.
- ▶ Separation of the **overall premium level** and the **structure of the portfolio**.

Multiplicative Tariff (3)

Notation for two rating factors \mathcal{I} and \mathcal{K} :

- ▶ tariff levels $i \in \{1, \dots, I\}$ for rating factor \mathcal{I}
- ▶ tariff levels $k \in \{1, \dots, K\}$ for rating factor \mathcal{K}
- ▶ tariff cells $(i, k) \in \{1, \dots, I\} \times \{1, \dots, K\} =: \mathcal{Q}$
- ▶ exposure $v_{i,k} \in (0, \infty)$ in tariff cell $(i, k) \in \mathcal{Q}$
- ▶ claim size $S_{i,k}$ in tariff cell (i, k)
- ▶ claim size per risk in tariff cell (i, k) :

$$\frac{S_{i,k}}{v_{i,k}}$$

- ▶ net premium per risk in tariff cell (i, k) :

$$E \left[\frac{S_{i,k}}{v_{i,k}} \right]$$

Multiplicative Tariff (4)

A collection of net premiums for the different cells is said to be a **multiplicative tariff** if there exist parameters

$\mu, \alpha_1, \dots, \alpha_I, \beta_1, \dots, \beta_K \in (0, \infty)$ such that the net premiums satisfy

$$E \left[\frac{S_{i,k}}{V_{i,k}} \right] = \mu \alpha_i \beta_k$$

for all $(i, k) \in Q$.

Possible scalings:

- ▶ $\max_{i=1}^I \alpha_i = 1 = \max_{k=1}^K \beta_k$
- ▶ $\sum_{i=1}^I \alpha_i = 1 = \sum_{k=1}^K \beta_k$

Notation:

- ▶ overall premium level: μ
- ▶ tariff factors: $\alpha_1, \dots, \alpha_I$ and β_1, \dots, β_K

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Collective Model (1)

Assume that

- ▶ $(\Omega, \mathcal{F}, \mathbb{P})$ is a probability space and that (M, \mathcal{M}) is a measurable space,
- ▶ N is a random variable $\Omega \rightarrow \mathbb{N}_0$, and
- ▶ $\{Y_j\}_{j \in \mathbb{N}}$ is a sequence of random variables $\Omega \rightarrow M$.

The pair

$$\langle N, \{Y_j\}_{j \in \mathbb{N}} \rangle$$

is said to be a **collective model** if the sequence $\{Y_j\}_{j \in \mathbb{N}}$ is independent and identically distributed and independent of N .

Interpretation:

- ▶ N is the **claim number** of a portfolio of risks
- ▶ Y_j is die **claim variable** of claim j of the portfolio, indicating the tariff cell, to which the risk causing this claim belongs, and the claim size.

Collective Model (2)

Aim: Decomposition of the collective model $\langle N, \{Y_j\}_{j \in \mathbb{N}} \rangle$ with respect to a measurable partition

$$M_1, \dots, M_m$$

of M with

$$\eta_p := \mathbf{P}[\{Y \in M_p\}]$$

and $\eta_p \in (0, 1)$ for all $p \in \{1, \dots, m\}$,
where Y denotes a random variable having the same distribution as each of the random variables Y_j .

To prepare the decomposition, we start with **thinning** of the collective model with respect to M_p for some $p \in \{1, \dots, m\}$.

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Thinning (1)

Thinning for fixed $p \in \{1, \dots, m\}$:

- ▶ Let

$$N_p := \sum_{j=1}^N \chi_{\{Y_j \in M_p\}}$$

denote the **number of claims** (for which the claim variable assumes a value) in M_p .

- ▶ Let $\nu_{p;0} := 0$ and for $j \in \mathbb{N}$ let

$$\nu_{p;j} := \inf \left\{ l \in \mathbb{N} \mid \nu_{p;j-1} < l \text{ und } Y_l \in M_p \right\}$$

denote the **number of the j -th claim** in M_p .

- ▶ For $j \in \mathbb{N}$, let

$$Y_{p;j} := \sum_{l=1}^{\infty} \chi_{\{\nu_{p;j}=l\}} Y_l$$

denote **claim variable of the j -th claim** in M_p .

Thinning (2)

Theorem. For every $p \in \{1, \dots, m\}$, the pair

$$\langle N_p, \{Y_{p;j}\}_{j \in \mathbb{N}} \rangle$$

is a collective model.

The independence of N_p and $\{Y_{p;j}\}_{j \in \mathbb{N}}$ is quite remarkable since each of these random variables has been constructed from the same sequence $\{Y_j\}_{j \in \mathbb{N}}$.

General application in statistics:

Post-stratifying does not destroy the property of sampling variables being independent and identically distributed.

It remains to investigate the dependencies between the thinned collective models.

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Decomposition (1)

Theorem. The families

$$\{Y_{1;j}\}_{j \in \mathbb{N}}, \dots, \{Y_{m;j}\}_{j \in \mathbb{N}}$$

are independent of each other.

Since $\langle N, \{Y_j\}_{j \in \mathbb{N}} \rangle$ is a collective model, each of

$$\langle N, \{\chi_{\{X_j \in M_p\}}\}_{j \in \mathbb{N}} \rangle$$

with $p \in \{1, \dots, m\}$ is a collective model as well.

In particular, **Wald's identities**

$$\begin{aligned} E[N_p] &= \eta_p E[N] \\ \text{var}[N_p] &= \eta_p E[N] + \eta_p^2 (\text{var}[N] - E[N]) \\ \text{cov}[N_p, N_q] &= \eta_p \eta_q (\text{var}[N] - E[N]) \end{aligned}$$

hold for all $p, q \in \{1, \dots, m\}$ with $p \neq q$.

Decomposition (2)

Theorem. The claim numbers satisfy

$$N = \sum_{p=1}^m N_p$$

and the identity

$$P \left[\bigcap_{p=1}^m \{N_p = n_p\} \mid \{N = n\} \right] = \frac{n!}{\prod_{p=1}^m n_p!} \prod_{p=1}^m \eta_p^{n_p}$$

holds for every $n \in \mathbb{N}_0$ and every family $\{n_p\}_{p \in \{1, \dots, m\}} \subseteq \mathbb{N}_0$ satisfying $\sum_{p=1}^m n_p = n$.

Theorem. The following are equivalent:

- ▶ The family $\{N_p\}_{p \in \{1, \dots, m\}}$ is independent.
- ▶ N has a Poisson distribution.

In this case, every claim number N_p has a Poisson distribution as well.

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Decomposition with Respect to Tariff Cells (1)

We consider the collective model $\langle N, \{Y_j\}_{j \in \mathbb{N}} \rangle$ with

$$M := (\{1, \dots, I\} \times \{1, \dots, K\}) \times \mathbb{R} = Q \times \mathbb{R}$$

and $\mathcal{M} := 2^{\{1, \dots, I\} \times \{1, \dots, K\}} \otimes \mathcal{B}(\mathbb{R}) = 2^Q \otimes \mathcal{B}(\mathbb{R})$. Then the individual claim variables can be represented in the form

$$Y_j = \begin{pmatrix} T_j \\ X_j \end{pmatrix}$$

where

- ▶ T_j denotes the **tariff cell** to which the risk causing claim j belongs, and
- ▶ X_j denotes the **claim size** of claim j .

We assume that, for every claim $j \in \mathbb{N}$, the random variables T_j and X_j are independent with $P[\{X_j \geq 0\}] = 1$.

Decomposition with Respect to Tariff Cells (2)

Since $\langle N, \{Y_j\}_{j \in \mathbb{N}} \rangle$ is a collective model,

$$\langle N, \{X_j\}_{j \in \mathbb{N}} \rangle$$

is a collective model as well.

Therefore, the **claim size** of the portfolio

$$S := \sum_{j=1}^N X_j$$

satisfies **Wald's identity**

$$E[S] = E[N] E[X]$$

where X denotes any random variable having the same distribution as each of the individual claim sizes X_j .

Decomposition with Respect to Tariff Cells (3)

Decomposition of the collective model $\langle N, \{Y_j\}_{j \in \mathbb{N}} \rangle$ with respect to the sets

$$M_{i,k} := \{(i, k)\} \times \mathbb{R}$$

with $\eta_{i,k} := P[\{Y \in M_{i,k}\}] = P[\{T = (i, k)\}]$ and $\eta_{i,k} \in (0, 1)$ yields the collective models

$$\langle N_{i,k}, \{Y_{i,k;j}\}_{j \in \mathbb{N}} \rangle$$

for the tariff cells. Moreover, the **individual claim sizes** of tariff cell (i, k) satisfy

$$Y_{i,k;j} = \begin{pmatrix} (i, k) \\ X_{i,k;j} \end{pmatrix}$$

and the **claim size** of tariff cell (i, k) satisfies

$$S_{i,k} := \sum_{j=1}^{N_{i,k}} X_{i,k;j}$$

Decomposition with Respect to Tariff Cells (4)

Theorem. The following identities hold for every tariff cell $(i, k) \in Q$:

- ▶ $P[\{X_{i,k} \in B\}] = P[\{X \in B\}]$ for all $B \in \mathcal{B}(\mathbb{R})$.
- ▶ $E[N_{i,k}] = \eta_{i,k} E[N]$.
- ▶ $E[S_{i,k}] = E[N_{i,k}] E[X_{i,k}] = \eta_{i,k} E[N] E[X] = \eta_{i,k} E[S]$.

Theorem.

- ▶ $N = \sum_{(i,k) \in Q} N_{i,k}$ and $S = \sum_{(i,k) \in Q} S_{i,k}$.
- ▶ The identity

$$P \left[\bigcap_{(i,k) \in Q} \{N_{i,k} = n_{i,k}\} \mid \{N = n\} \right] = \frac{n!}{\prod_{(i,k) \in Q} n_{i,k}!} \prod_{(i,k) \in Q} \eta_{i,k}^{n_{i,k}}$$

holds for all $n \in \mathbb{N}_0$ and every family $\{n_{i,k}\}_{(i,k) \in Q} \subseteq \mathbb{N}_0$ satisfying $\sum_{(i,k) \in Q} n_{i,k} = n$ (conditional multinomial distribution).

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Assumptions (1)

Given: Known measures of exposure

$$v_{i,k} \in \mathbb{R}_+$$

for the tariff cells $(i, k) \in Q$.

Assumption: There exist parameters $\alpha_1, \dots, \alpha_I \in (0, \infty)$ and $\beta_1, \dots, \beta_K \in (0, \infty)$ such that

$$\eta_{i,k} = \alpha_i \beta_k v_{i,k}$$

holds for all $(i, k) \in Q$ (threefold proportionality).

The assumption implies

$$\sum_{(i,k) \in Q} \alpha_i \beta_k v_{i,k} = 1$$

Assumptions (2)

Because of the assumption $\eta_{i,k} = \alpha_i \beta_k v_{i,k}$,

- ▶ the expected **claim number** in tariff cell (i, k) satisfies

$$E[N_{i,k}] = \eta_{i,k} E[N] = \alpha_i \beta_k v_{i,k} E[N]$$

- ▶ the expected **claim size** in tariff cell (i, k) satisfies

$$E[S_{i,k}] = \eta_{i,k} E[S] = \alpha_i \beta_k v_{i,k} E[S]$$

- ▶ the **net premium per risk** in tariff cell (i, k) satisfies

$$E[S_{i,k}/v_{i,k}] = \alpha_i \beta_k E[S]$$

In particular, the assumption yields a multiplicative tariff.

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Marginal–Sum Estimation: Claim Sizes per Cell (1)

The **claim sizes** of the tariff cells and the total claim size of the portfolio are related by the identities

$$\sum_{(i,k) \in Q} S_{i,k} = S$$

and

$$E[S_{i,k}] = \alpha_i \beta_k v_{i,k} E[S]$$

Marginal–Sum Estimation: Claim Sizes per Cell (2)

Summation over the equations $E[S_{i,k}] = \alpha_i \beta_k v_{i,k} E[S]$ yields

$$\begin{aligned} \sum_{k=1}^K E[S_{i,k}] &= \sum_{k=1}^K \alpha_i \beta_k v_{i,k} E[S] & i \in \{1, \dots, I\} \\ \sum_{i=1}^I E[S_{i,k}] &= \sum_{i=1}^I \alpha_i \beta_k v_{i,k} E[S] & k \in \{1, \dots, K\} \end{aligned}$$

Estimators $\hat{\alpha}_1, \dots, \hat{\alpha}_I$ and $\hat{\beta}_1, \dots, \hat{\beta}_K$ of the parameters $\alpha_1, \dots, \alpha_I$ and β_1, \dots, β_K are said to be **Marginal–Sum estimators** if they satisfy the **Marginal–Sum equations**

$$\begin{aligned} \sum_{k=1}^K \hat{\alpha}_i \hat{\beta}_k v_{i,k} &= \sum_{k=1}^K \frac{S_{i,k}}{S} & i \in \{1, \dots, I\} \\ \sum_{i=1}^I \hat{\alpha}_i \hat{\beta}_k v_{i,k} &= \sum_{i=1}^I \frac{S_{i,k}}{S} & k \in \{1, \dots, K\} \end{aligned}$$

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Marginal–Sum Estimation: Claim Numbers (1)

The **claim numbers** of the tariff cells and the total claim number of the portfolio are related by the identities

$$\sum_{(i,k) \in Q} N_{i,k} = N$$

and

$$E[N_{i,k}] = \alpha_i \beta_k v_{i,k} E[N]$$

Note that

$$E[S] = E[N] E[X]$$

which because of $E[X_{i,k}] = E[X]$ implies

$$E[S_{i,k}] = E[N_{i,k}] E[X]$$

Marginal–Sum Estimation: Claim Numbers (2)

Summation over the equations $E[N_{i,k}] = \alpha_i \beta_k v_{i,k} E[N]$ yields

$$\sum_{k=1}^K E[N_{i,k}] = \sum_{k=1}^K \alpha_i \beta_k v_{i,k} E[N] \quad i \in \{1, \dots, I\}$$

$$\sum_{i=1}^I E[N_{i,k}] = \sum_{i=1}^I \alpha_i \beta_k v_{i,k} E[N] \quad k \in \{1, \dots, K\}$$

and hence the **Marginal–Sum equations**

$$\sum_{k=1}^K \hat{\alpha}_i \hat{\beta}_k v_{i,k} = \sum_{k=1}^K \frac{N_{i,k}}{N} \quad i \in \{1, \dots, I\}$$

$$\sum_{i=1}^I \hat{\alpha}_i \hat{\beta}_k v_{i,k} = \sum_{i=1}^I \frac{N_{i,k}}{N} \quad k \in \{1, \dots, K\}$$

(now with $N_{i,k}/N$ in the place of $S_{i,k}/S$).

Marginal–Sum Estimation: Claim Numbers (3)

In the case where $v_{i,k} > 0$ holds for all (i, k) ,
Marginal–Sum estimation based on claim numbers can be
justified by the **Maximum–Likelihood principle**:

Define $p : \mathbb{N}_0 \rightarrow \mathbb{R}_+$ by letting

$$p(n) := P[\{N = n\}]$$

Then every family $\{n_{i,k}\}_{(i,k) \in Q} \subseteq \mathbb{N}_0$ with $n := \sum_{(i,k) \in Q} n_{i,k}$
satisfies

$$\begin{aligned} P \left[\bigcap_{(i,k) \in Q} \{N_{i,k} = n_{i,k}\} \right] &= \frac{n!}{\prod_{(i,k) \in Q} n_{i,k}!} \prod_{(i,k) \in Q} \eta_{i,k}^{n_{i,k}} \cdot p(n) \\ &= \frac{n!}{\prod_{(i,k) \in Q} n_{i,k}!} \prod_{(i,k) \in Q} (\alpha_i \beta_k v_{i,k})^{n_{i,k}} \cdot p(n) \end{aligned}$$

Marginal–Sum Estimation: Claim Numbers (4)

Likelihood Function:

$$\begin{aligned}
 L(\hat{\alpha}_1, \dots, \hat{\alpha}_I, \hat{\beta}_1, \dots, \hat{\beta}_K \mid \{N_{i,k}\}_{(i,k) \in Q}) \\
 = \frac{N!}{\prod_{(i,k) \in Q} N_{i,k}!} \prod_{(i,k) \in Q} (\hat{\alpha}_i \hat{\beta}_k v_{i,k})^{N_{i,k}} \cdot p(N)
 \end{aligned}$$

Log–Likelihood Function:

$$\begin{aligned}
 (\log \circ L)(\hat{\alpha}_1, \dots, \hat{\alpha}_I, \hat{\beta}_1, \dots, \hat{\beta}_K \mid \{N_{i,k}\}_{(i,k) \in Q}) \\
 = \sum_{(i,k) \in Q} N_{i,k} \log(\hat{\alpha}_i \hat{\beta}_k v_{i,k}) + C
 \end{aligned}$$

Marginal–Sum Estimation: Claim Numbers (5)

Lagrange–Approach: Minimize

$$\begin{aligned} & \ell\left(\hat{\alpha}_1, \dots, \hat{\alpha}_I, \hat{\beta}_1, \dots, \hat{\beta}_K, \lambda \mid \{N_{i,k}\}_{(i,k) \in Q}\right) \\ &= \sum_{(i,k) \in Q} N_{i,k} \log(\hat{\alpha}_i \hat{\beta}_k v_{i,k}) + \lambda \left(1 - \sum_{(i,k) \in Q} \hat{\alpha}_i \hat{\beta}_k v_{i,k}\right) \end{aligned}$$

Partial differentiation yields

$$\begin{aligned} \sum_{k=1}^K \frac{N_{i,k}}{\hat{\alpha}_i} &= \lambda \sum_{k=1}^K \hat{\beta}_k v_{i,k} & i \in \{1, \dots, I\} \\ \sum_{i=1}^I \frac{N_{i,k}}{\hat{\beta}_k} &= \lambda \sum_{i=1}^I \hat{\alpha}_i v_{i,k} & k \in \{1, \dots, K\} \\ 1 &= \sum_{(i,k) \in Q} \hat{\alpha}_i \hat{\beta}_k v_{i,k} \end{aligned}$$

Marginal–Sum Estimation: Claim Numbers (6)

and hence

$$\begin{aligned} \sum_{k=1}^K N_{i,k} &= \lambda \sum_{k=1}^K \hat{\alpha}_i \hat{\beta}_k v_{i,k} & i \in \{1, \dots, I\} \\ \sum_{i=1}^I N_{i,k} &= \lambda \sum_{i=1}^I \hat{\alpha}_i \hat{\beta}_k v_{i,k} & k \in \{1, \dots, K\} \\ 1 &= \sum_{(i,k) \in Q} \hat{\alpha}_i \hat{\beta}_k v_{i,k} \end{aligned}$$

Since

$$N = \sum_{(i,k) \in Q} N_{i,k} = \lambda \sum_{(i,k) \in Q} \hat{\alpha}_i \hat{\beta}_k v_{i,k} = \lambda$$

this yields the Marginal–Sum equations

$$\begin{aligned} \sum_{k=1}^K \hat{\alpha}_i \hat{\beta}_k v_{i,k} &= \sum_{k=1}^K \frac{N_{i,k}}{N} & i \in \{1, \dots, I\} \\ \sum_{i=1}^I \hat{\alpha}_i \hat{\beta}_k v_{i,k} &= \sum_{i=1}^I \frac{N_{i,k}}{N} & k \in \{1, \dots, K\} \end{aligned}$$

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Marginal–Sum Problem (1)

Given: Matrices $\mathbf{V}, \mathbf{S} \in \mathbb{R}_+^{\{1, \dots, I\} \times \{1, \dots, K\}}$ with the property, that every row and every column contains at least one element $\neq 0$:

Wanted: Solution(s) $(\mu, \alpha, \beta) \in (0, \infty) \times (0, \infty)^I \times (0, \infty)^K$ to the **Marginal–Sum equations (MS equations)**

$$\begin{aligned} \sum_{k=1}^K \mu \alpha_i \beta_k v_{i,k} &= \sum_{k=1}^K s_{i,k} & i \in \{1, \dots, I\} \\ \sum_{i=1}^I \mu \alpha_i \beta_k v_{i,k} &= \sum_{i=1}^I s_{i,k} & k \in \{1, \dots, K\} \end{aligned}$$

Remark: If (μ, α, β) is a solution, then μ satisfies

$$\mu = \frac{\sum_{i=1}^I \sum_{k=1}^K s_{i,k}}{\sum_{i=1}^I \sum_{k=1}^K \alpha_i v_{i,k} \beta_k} = \frac{\mathbf{1}' \mathbf{S} \mathbf{1}}{\alpha' \mathbf{V} \beta}$$

Therefore, only (α, β) will be referred to as a solution.

Marginal-Sum Problem (2)

The MS equations have a **radially unique solution** if

- ▶ they have a solution (α^*, β^*) and
- ▶ every solution (α, β) has the form $(c\alpha^*, d\beta^*)$ for some $c, d \in (0, \infty)$.

Example. Assume that $v_{i,k} = v > 0$ holds for all (i, k) . Then the MS equations become

$$(\mu v \mathbf{1}'\beta) \alpha = \mathbf{S}\mathbf{1}$$

$$(\mu v \mathbf{1}'\alpha) \beta = \mathbf{S}'\mathbf{1}$$

and hence have the (radially unique) solution (α^*, β^*) with

$$\alpha^* = \mathbf{S}\mathbf{1}$$

$$\beta^* = \mathbf{S}'\mathbf{1}$$

and $\mu^* = (v \mathbf{1}'\mathbf{S}\mathbf{1})^{-1}$.

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Equivalent Fixed-Point Problems (1)

Define $G : (0, \infty)^I \rightarrow (0, \infty)^K$ and $H : (0, \infty)^K \rightarrow (0, \infty)^I$ by letting

$$G_k(\alpha) := \frac{\mathbf{1}'\mathbf{S}\mathbf{e}_k}{\alpha'\mathbf{V}\mathbf{e}_k} \quad \text{und} \quad H_i(\beta) := \frac{\mathbf{e}_i'\mathbf{S}\mathbf{1}}{\mathbf{e}_i'\mathbf{V}\beta}$$

Then:

- ▶ G and H are continuous, decreasing, and homogeneous of degree -1 .
- ▶ (α, β) is a solution to the MS equations iff there is some $\mu \in (0, \infty)$ such that

$$\mu \alpha = H(\beta) \quad \text{and} \quad \mu \beta = G(\alpha)$$

- ▶ If (α, β) is a solution to the MS equations, then

$$\mu = \frac{\mathbf{1}'\mathbf{S}\mathbf{1}}{\alpha'\mathbf{V}\beta}$$

Equivalent Fixed-Point Problem (2)

Define $\Phi : (0, \infty)^I \rightarrow (0, \infty)^I$ and $\Psi : (0, \infty)^K \rightarrow (0, \infty)^K$ by letting

$$\Phi := H \circ G \quad \text{and} \quad \Psi := G \circ H$$

Then Φ and Ψ are continuous, increasing, and homogeneous of degree 1 and $G \circ \Phi = \Psi \circ G$ and $\Phi \circ H = H \circ \Psi$. **Thus:** Every strictly positive multiple of a fixed point is a fixed point as well.

Theorem.

- ▶ If (α, β) is a solution to the MS equations, then α is a fixed point of Φ and β is a fixed point of Ψ .
- ▶ If α is a fixed point of Φ , then $G(\alpha)$ is a fixed point of Ψ and $(\alpha, G(\alpha))$ is a solution to the MS equations.
- ▶ If β is a fixed point of Ψ , then $H(\beta)$ is a fixed point of Φ and $(H(\beta), \beta)$ is a solution to the MS equations.

Equivalent Fixed-Point Problem (3)

Proof.

- ▶ Assume that (α, β) is a solution to the MS equations and let $\mu := \mathbf{1}'\mathbf{S}\mathbf{1}/\alpha'\mathbf{V}\beta$. Then

$$\mu \alpha = H(\beta) = H(\mu^{-1} G(\alpha)) = \mu H(G(\alpha)) = \mu \Phi(\alpha)$$

In particular, α is a fixed point of Φ .

- ▶ Assume that α is a fixed point of Φ . Then

$$\Psi(G(\alpha)) = G(\Phi(\alpha)) = G(\alpha)$$

In particular,

$$\beta := G(\alpha)$$

is a fixed point of Ψ and satisfies

$$\alpha = \Phi(\alpha) = H(G(\alpha)) = H(\beta)$$

Therefore, (α, β) is a solution to the MS equations.

Equivalent Fixed-Point Problem (4)

Theorem. The following are equivalent:

- ▶ The MS equations have a radially unique solution.
- ▶ Φ has a radially unique fixed point.
- ▶ Ψ has a radially unique fixed point.

Existence Theorem. Assume that $v_{i,k} > 0$ holds for all (i, k) . Then Φ has a fixed point.

Sketch of the proof: Define $\bar{G} : \mathbb{R}_+^I \setminus \{\mathbf{0}\} \rightarrow (0, \infty)^K$ by letting

$$\bar{G}_k(\alpha) := \frac{\mathbf{1}' \mathbf{S} \mathbf{e}_k}{\alpha' \mathbf{V} \mathbf{e}_k}$$

and define

$$\bar{\Phi} := H \circ \bar{G}$$

Equivalent Fixed-Point Problem (5)

- ▶ \bar{G} and $\bar{\Phi}$ are continuous extensions of G and Φ .
- ▶ Let

$$\Delta' := \left\{ \alpha \in \mathbb{R}'_+ \mid \mathbf{1}'\alpha = 1 \right\}$$

and define $\tilde{\Phi} : \Delta' \rightarrow \Delta'$ by letting

$$\tilde{\Phi}(\alpha) := \frac{1}{1 + \mathbf{1}'\bar{\Phi}(\alpha)} \left(\alpha + \bar{\Phi}(\alpha) \right)$$

Then Δ' is nonempty, convex and compact, and $\tilde{\Phi}$ is continuous.

- ▶ Brouwer's Fixed Point Theorem: $\tilde{\Phi}$ has a fixed point α .
- ▶ α is also a fixed point of $\bar{\Phi}$ (use definition of $\bar{\Phi}$).
- ▶ Since $\alpha = \bar{\Phi}(\alpha) \in (0, \infty)'$, α is also a fixed point of Φ .

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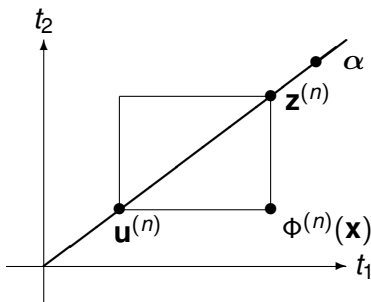
Radial Uniqueness and Iteration (1)

Theorem. Assume that $v_{i,k} > 0$ holds for all (i, k) .

If α is a fixed point of Φ , then, for every $\mathbf{x} \in (0, \infty)^l$, there exists some $\xi \in (0, \infty)$ such that

$$\lim_{n \rightarrow \infty} \Phi^n(\mathbf{x}) = \xi \alpha$$

Sketch of the proof:



Radial Uniqueness and Iteration (2)

Corollary. Assume that $v_{i,k} > 0$ holds for all (i, k) .
Then Φ has a radially unique fixed point.

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Example (1)

Example. Let

$$\mathbf{v} := \begin{pmatrix} 2 & 0 \\ 3 & 1 \end{pmatrix} \quad \text{und} \quad \mathbf{s} := \begin{pmatrix} 1 & 0 \\ 0 & 4 \end{pmatrix}$$

Then we have $v_{1,2} = 0$ (!) and

$$2\mu\alpha_1\beta_1 = 1$$

$$3\mu\alpha_2\beta_1 + \mu\alpha_2\beta_2 = 4$$

$$2\mu\beta_1\alpha_1 + 3\mu\beta_1\alpha_2 = 1$$

$$\mu\beta_2\alpha_2 = 4$$

Therefore, the MS equations have no solution.

Example (2)

We have

$$\Phi(\alpha) = \begin{pmatrix} 1 \\ \frac{2/(2\alpha_1 + 3\alpha_2)}{4} \\ \frac{3/(2\alpha_1 + 3\alpha_2) + 4/\alpha_2}{4} \end{pmatrix}$$

Iteration with the initial value $\mathbf{x} := (1/2) \mathbf{1}$ yields the following scaled iterates $\mathbf{x}^{(n)} := \Phi^{(n)}(\mathbf{x}) / \mathbf{1}'\Phi^{(n)}(\mathbf{x})$:

n	$x_1^{(n)}$	$x_2^{(n)}$
0	0,5	0,5
1	0,742	0,258
10	0,952	0,048
100	0,995	0,005
1000	0,999	0,001
10000	1,000	0,000

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DSVM = Dresdner Schriften zur Versicherungsmathematik.